

# Maximum Likelihood Density Estimation under Total Positivity

Elina Robeva  
MIT

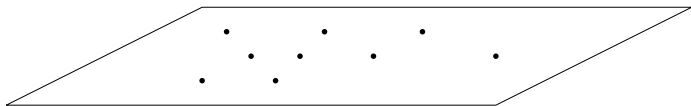
joint work with Bernd Sturmfels, Ngoc Tran, and Caroline Uhler  
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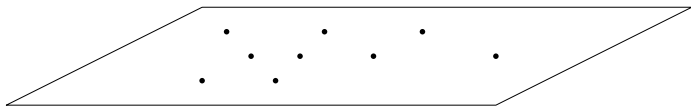
# Density estimation

Given i.i.d. samples  $X = \{x_1, \dots, x_n\} \subset \mathbb{R}^d$  from an unknown distribution on  $\mathbb{R}^d$  with density  $p$ , can we estimate  $p$ ?



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- parametric: assume that  $p$  lies in some parametric family, and estimate parameters
  - finite-dimensional problem
  - too restrictive; the real-world distribution might not lie in the specified parametric family
- non-parametric: assume that  $p$  lies in a non-parametric family, e.g. impose shape-constraints on  $p$  (convex, log-concave, monotone, etc.)
  - infinite-dimensional problem
  - need constraints that are:
    - strong enough so that there is no spiky behavior
    - weak enough so that function class is large

# Shape-constrained density estimation

- monotonically decreasing densities: [Grenander 1956, Rao 1969]
- convex densities: [Anevski 1994, Groeneboom, Jongbloed, and Wellner 2001]
- **log-concave** densities: [Cule, Samworth, and Stewart 2008]
- generalized additive models with shape constraints: [Chen and Samworth 2016]
- this talk: **totally positive and log-concave** densities

## MTP<sub>2</sub> distributions

- A distribution with density  $p$  on  $\mathcal{X} \subseteq \mathbb{R}^d$  is *multivariate totally positive of order 2* (or  $MTP_2$ ) if

$$p(x)p(y) \leq p(x \wedge y)p(x \vee y) \quad \text{for all } x, y \in \mathcal{X},$$

where  $x \wedge y$  and  $x \vee y$  are the componentwise minimum and maximum.

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- MTP<sub>2</sub> is the same as *log-supermodular*:

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- A random vector  $X$  taking values in  $\mathbb{R}^d$  is *positively associated* if for any non-decreasing functions  $\phi, \psi : \mathbb{R}^d \rightarrow \mathbb{R}$

$$\text{cov}(\phi(X), \psi(X)) \geq 0.$$

- MTP<sub>2</sub> implies positive association (Fortuin Kasteleyn Ginibre inequality, 1971).

# Properties of $MTP_2$ distributions

Theorem (F<sub>allat</sub>, L<sub>auritzen</sub>, S<sub>adeghi</sub>, U<sub>hler</sub>, W<sub>ermuth</sub> and Z<sub>wiernik</sub>, 2015)

If  $X = (X_1, \dots, X_d)$  is  $MTP_2$ , then

- (i) any marginal distribution is  $MTP_2$ ,
- (ii) any conditional distribution is  $MTP_2$ ,
- (iii)  $X$  has the marginal independence structure

$$X_i \perp\!\!\!\perp X_j \iff \text{cov}(X_i, X_j) = 0.$$

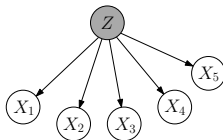
Theorem (K<sub>arlin</sub> and R<sub>inott</sub>, 1980)

If  $p(x) > 0$  and  $p$  is  $MTP_2$  for any pair of coordinates when the others are held constant, then  $p$  is  $MTP_2$ .



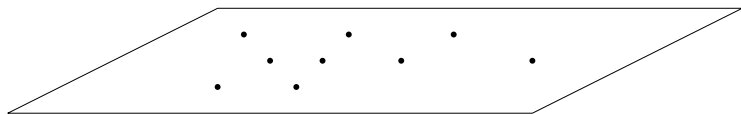
## Examples of MTP<sub>2</sub> distributions

- A Gaussian random variable  $X \sim \mathcal{N}(\mu, \Sigma)$  is MTP<sub>2</sub> whenever  $\Sigma^{-1}$  is an M-matrix, i.e. its off-diagonal entries are nonpositive.
- The joint distribution of observed variables influenced by one hidden variable



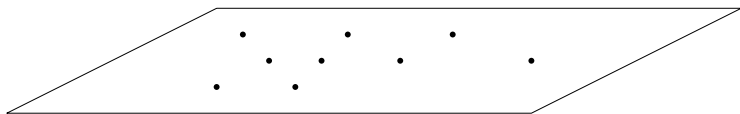
- Very common in real data: e.g. IQ test scores, phylogenetics data, financial econometrics data, and others
- Many models imply MTP<sub>2</sub>:
  - Ferromagnetic Ising models
  - Order statistics of i.i.d. variables
  - Brownian motion tree models
  - Latent tree models (e.g. single factor analysis models)

# Maximum Likelihood Estimation



Given i.i.d. samples  $X = \{x_1, \dots, x_n\} \subset \mathbb{R}^d$  with weights  $w = (w_1, \dots, w_n)$  (where  $w_1, \dots, w_n \geq 0$ ,  $\sum w_i = 1$ ) from a distribution  $p$  on  $\mathbb{R}^d$ , can we estimate  $p$ ?

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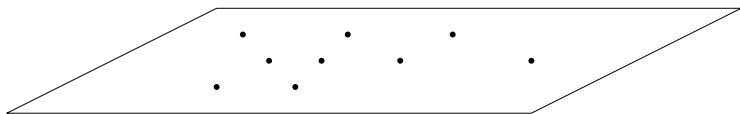


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The *log-likelihood* of observing  $X = \{x_1, \dots, x_n\}$  with weights  $w = (w_1, \dots, w_n)$  if they are drawn i.i.d. from  $p$  is (up to an additive constant)

$$\ell_p(X, w) := \sum_{i=1}^n w_i \log(p(x_i)).$$

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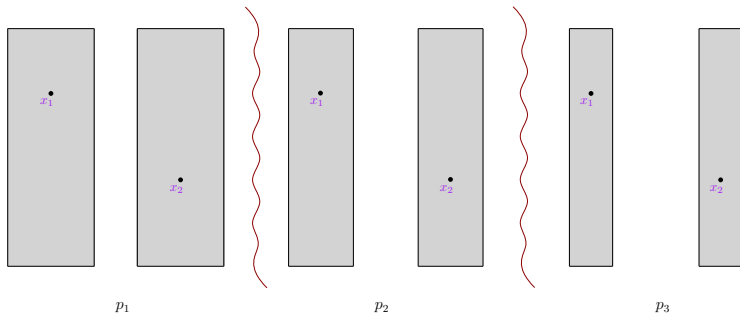
We would like to

$$\begin{aligned} & \text{maximize}_p \quad \sum_{i=1}^n w_i \log(p(x_i)) \\ & \text{s.t.} \quad p \text{ is an MTP}_2 \text{ density.} \end{aligned}$$

# Maximum Likelihood Estimation under MTP<sub>2</sub>

Suppose we observe two points:  $X = \{x_1, x_2\} \subset \mathbb{R}^2$ . We can find a sequence of MTP<sub>2</sub> densities  $p_1, p_2, p_3, \dots$  such that

$$\ell_{p_n}(X) \rightarrow \infty \quad \text{as } n \rightarrow \infty.$$

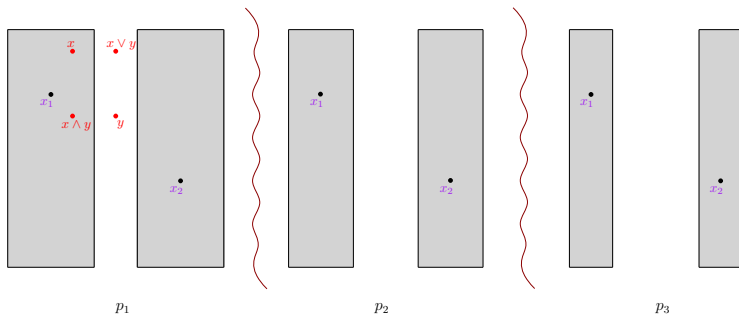


Thus, the MLE doesn't exist.

# Maximum Likelihood Estimation under $\text{MTP}_2$

Suppose we observe two points:  $X = \{x_1, x_2\} \subset \mathbb{R}^2$ . We can find a sequence of  $\text{MTP}_2$  densities  $p_1, p_2, p_3, \dots$  such that

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Thus, the MLE doesn't exist.

## Maximum Likelihood Estimation under $\text{MTP}_2$

To ensure that the likelihood function is bounded, we impose the condition that  $p$  is log-concave.

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A function  $f : \mathbb{R}^d \rightarrow \mathbb{R}$  is *log-concave* if its logarithm is concave.



## Maximum Likelihood Estimation under $MTP_2$

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A function  $f : \mathbb{R}^d \rightarrow \mathbb{R}$  is *log-concave* if its logarithm is concave.

- Log-concavity is a natural assumption because it ensures the density is continuous and includes many known families of parametric distributions.
- Log-concave families:
  - Gaussian; Uniform( $a, b$ ); Gamma( $k, \theta$ ) for  $k \geq 1$ ; Beta( $a, b$ ) for  $a, b \geq 1$ .
- Maximum likelihood estimation under log-concavity is a well-studied problem (Cule et al. 2008, Dümbgen et al. 2009, Schuhmacher et al. 2010, ...).

# Maximum Likelihood Estimation under Log-Concavity

$$\begin{array}{ll} \text{maximize}_p & \sum_{i=1}^n w_i \log(p(x_i)) \\ \text{s.t.} & p \text{ is a density} \\ \text{and} & p \text{ is log-concave.} \end{array}$$

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## Theorem (Cule, Samworth and Stewart 2008)

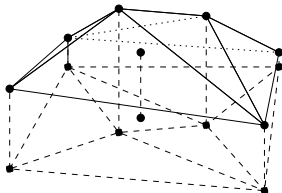
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# Maximum Likelihood Estimation under Log-Concavity

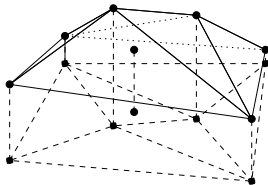
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## Theorem (Cule, Samworth and Stewart 2008)

- *With probability 1, a log-concave maximum likelihood estimator  $\hat{p}$  exists and is unique.*
- *Moreover,  $\log(\hat{p})$  is a 'tent-function' supported on the convex hull of the data  $P(X) = \text{conv}(x_1, \dots, x_n)$ .*



# Optimizing over Tent Functions

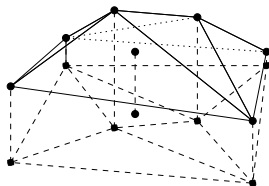


Given points  $X = \{x_1, \dots, x_n\}$  and heights  $y = (y_1, \dots, y_n) \in \mathbb{R}^n$ , the *tent function*

$$h_{X,y} : \mathbb{R}^d \rightarrow \mathbb{R}$$

is the smallest concave function such that  $h_{X,y}(x_i) \geq y_i$  for all  $i$ . Thus,  $\hat{p} = \exp(h_{X,y})$  for some  $y$ .

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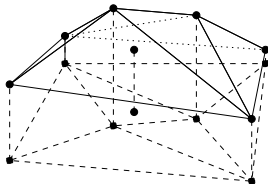
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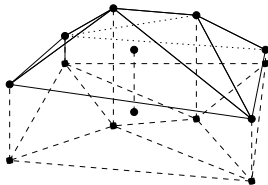
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INFINITE DIMENSIONAL

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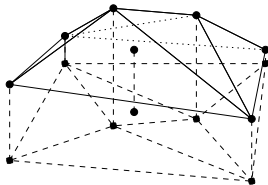
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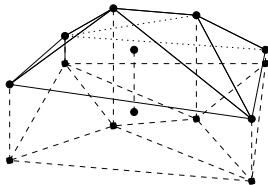
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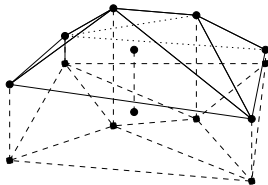
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$$\max_{y \in \mathbb{R}^n} \quad \sum_{i=1}^n w_i y_i - \int \exp(h_{X,y}(t)) dt$$

# Maximum Likelihood Estimation under Log-concavity and $MTP_2$

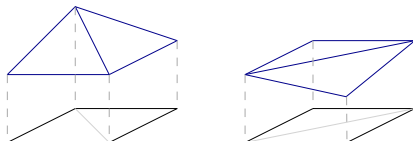
Questions:

1. Does the MLE under log-concavity and  $MTP_2$  exist with probability 1 and, if so, is it unique?
2. What is the shape of the MLE under log-concavity and  $MTP_2$ ?
  - 2.1 What is the support of the MLE?
  - 2.2 Is the MLE always  $\exp(\text{tent function})$ ?
3. Which tent functions are allowed?
4. Can we compute the MLE?

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Recall:  $p$  is  $MTP_2$  if and only if  $\log(p)$  is *supermodular*, i.e.

$$\log p(x) + \log p(y) \leq \log p(x \wedge y) + \log p(x \vee y), \text{ for all } x, y.$$

# Existence and Uniqueness of the MLE

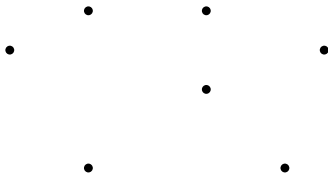
Theorem (R., Sturmfels, Tran, Uhler)

*The maximum likelihood estimator under log-concavity and  $MTP_2$  exists and is unique with probability 1 as long as there are at least 3 samples.*

Proof uses convergence properties for log-concave distributions, and does not shed light on the shape of the MLE.

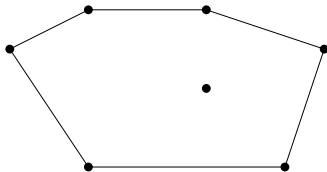
# The Support of the MLE

Consider the following samples:



# The Support of the MLE

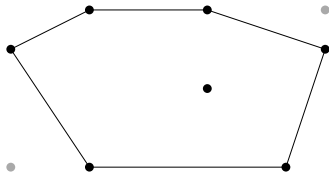
Under log-concavity, the support of the MLE is the convex hull:





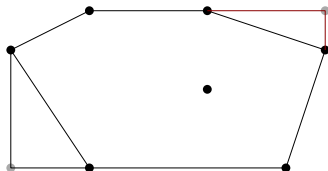
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Under log-concavity and  $\text{MTP}_2$  we need the density to be nonzero at more points:



# The Support of the MLE

Under log-concavity and  $\text{MTP}_2$  we need the density to be nonzero at more points:



and we need the convex hull of all of these points.

Support of the MLE = "min-max convex hull" of  $X$ .

# The Min-Max Convex Hull

## Definition

$MM(X)$  = smallest *min-max closed* set  $S$  containing  $X$ , i.e.  $x, y \in S \Rightarrow x \wedge y, x \vee y \in S$

$MMconv(X)$  = smallest *min-max closed and convex* set containing  $X$ .

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- How can we find  $\text{MMconv}(X)$  for  $X = \{x_1, \dots, x_n\} \subseteq \mathbb{R}^d$ ?

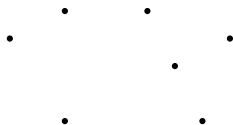
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- Intuitive first proposal:



Start with  $X$ .

Add points to  $X$   
until we get  $MM(X)$ .

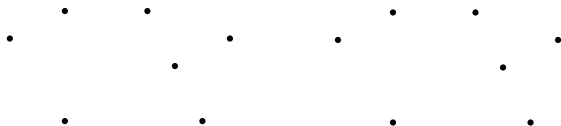
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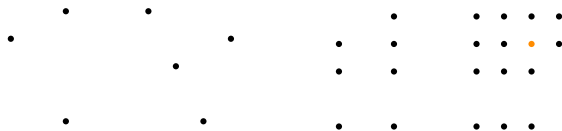
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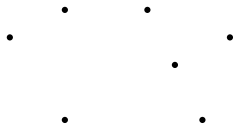
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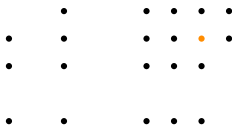
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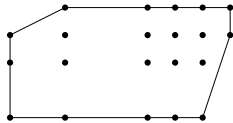
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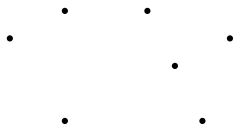
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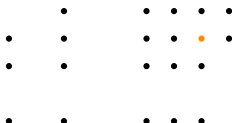
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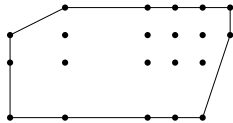
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- Is it always true that  $MMconv(X) = \text{conv}(MM(X))$ ?

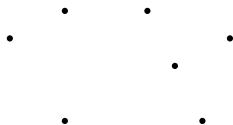
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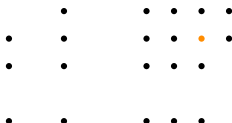
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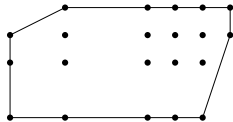
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Take  $\text{conv}(MM(X))$ .

- Is it always true that  $MMconv(X) = \text{conv}(MM(X))$ ? No!

# The Min-Max Convex Hull

## Lemma

Let  $X = \{x_1, \dots, x_n\}$ . If  $X \subseteq \mathbb{R}^2$  or  $X \subseteq \{0, 1\}^d$ , then,

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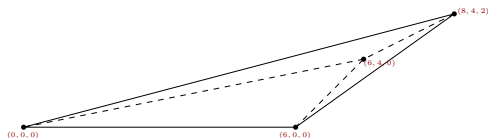
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It turns out that

$$MM(X) = X.$$

But

$conv(MM(X))$  is **not min-max closed!**

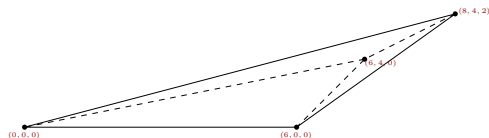
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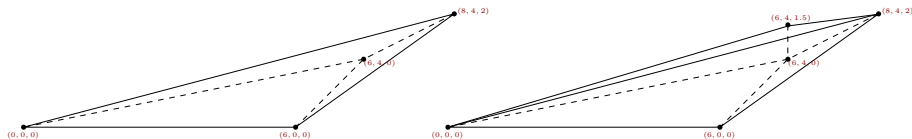
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Therefore,

$$conv(MM(X)) \subsetneq MMconv(X).$$

# The 2-D Projections Theorem

## Theorem (The 2-D Projections Theorem)

For any finite subset  $X \subseteq \mathbb{R}^d$ . Then we have

$$MMconv(X) = \bigcap_{1 \leq i < j \leq d} \pi_{ij}^{-1}(\text{conv}(\pi_{ij}(MM(X)))) .$$

$$\begin{aligned} \pi_{ij} : \mathbb{R}^d &\rightarrow \mathbb{R}, \\ x &\mapsto (x_i, x_j). \end{aligned}$$



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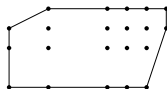
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## Corollary (Queyranne and Tardella, 2006)

A subset  $C$  in  $\mathbb{R}^d$  is a min-max closed convex polytope if and only if it is defined by a finite collection of bimonotone linear inequalities.

A linear inequality  $a \cdot x + b \leq 0$  is *bimonotone* if it has the form

$$a_i x_i + a_j x_j + b \leq 0, \quad \text{where} \quad a_i a_j \leq 0.$$



# Back to Log-concave and $MTP_2$ Maximum Likelihood Estimation

1. Does the MLE under log-concavity and  $MTP_2$  exist with probability 1 and, if so, is it unique? **Yes.**
2. What is the shape of the MLE under log-concavity and  $MTP_2$ ?
  - 2.1 What is the support of the MLE?  **$MMconv(X)$ ; We can compute it.**
  - 2.2 Is the MLE always exp(tent function)?
3. Which tent functions are allowed?
4. Can we compute the MLE?

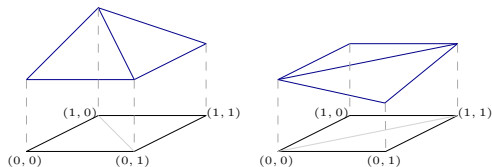
# Supermodular Tent Functions

Recall that  $p = \exp(h)$  is **MTP**<sub>2</sub> if and only if  $h$  is **supermodular**, i.e.

$$h(x) + h(y) \leq h(x \wedge y) + h(x \vee y), \quad \text{for all } x, y \in \mathbb{R}^d.$$

## Theorem (R., Sturmfels, Tran, Uhler)

Let  $X \subset \mathbb{R}^d$  be a finite set of points. A tent function  $h$  is supermodular if and only if all of the walls of the subdivision  $h$  induces are **bimonotone**.



## Remark

If we want to find the best supermodular  $h_{X,y}$ , we need to optimize over the set of heights  $y$  that induce bimonotone subdivisions.

- In general not convex.
- Example:  $X = \{0, 1\} \times \{0, 1\} \times \{0, 1, 2\}$ .

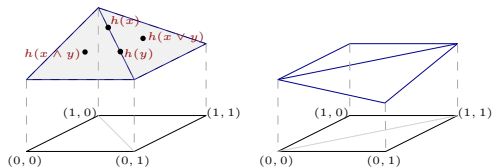
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# Is the MLE is the exponential of a tent function?

1. Does the MLE under log-concavity and  $\text{MTP}_2$  exist with probability 1 and, if so, is it unique? **Yes.**
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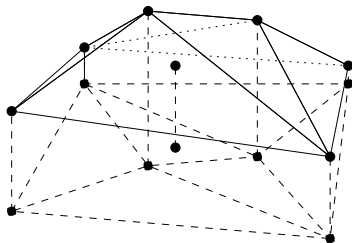
# Why is the Log-concave MLE the exponential of a tent function?

Recall:

$$\begin{array}{ll} \text{maximize}_p & \sum_{i=1}^n w_i \log(p(x_i)) \\ \text{s.t.} & p \text{ is a density} \\ \text{and} & p \text{ is log-concave.} \end{array}$$

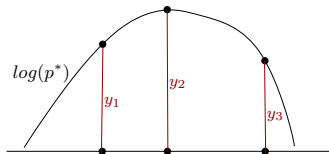
## Theorem (Cule, Samworth and Stewart 2008)

- With probability 1, a log-concave maximum likelihood estimator  $\hat{p}$  exists and is unique.
- Moreover,  $\log(\hat{p})$  is a 'tent-function' supported on the convex hull of the data  $P(X) = \text{conv}(x_1, \dots, x_n)$ .



# Why is the Log-concave MLE the exponential of a tent function?

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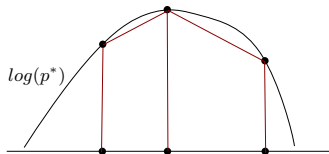


Proof of theorem:

- Suppose that  $p^*$  is the MLE and that  $\log p^*$  is not a tent function.
- Let  $y_i = \log p^*(x_i)$ ,  $i = 1, \dots, n$ .
- Consider  $p = \exp(h_{X,y})$ . It gives a higher objective value than  $p^*$ .
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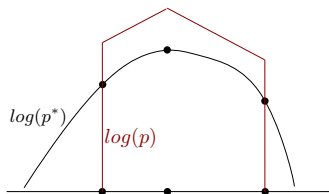
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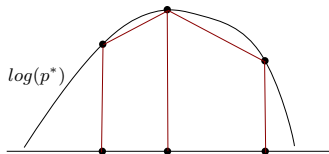


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# Proving that the Log-concave MTP<sub>2</sub> MLE is the exponential of a tent function

$$\begin{aligned} & \text{maximize}_p \quad \sum_{i=1}^n w_i \log(p(x_i)) \\ & \text{s.t.} \quad p \text{ is a log-concave density} \\ & \text{and} \quad p \text{ is MTP}_2. \end{aligned}$$

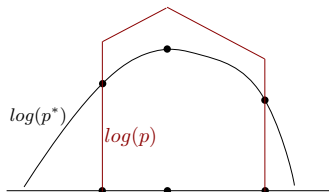


Proof that the MLE is a tent function:

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- Suppose that  $p^*$  is the MLE and that  $\log p^*$  is not a tent function
- Let  $y_i = \log p^*(x_i)$ ,  $i = 1, \dots, n$ .
- Consider  $p = \exp(h_{X,Y})$ . It gives a higher objective value than  $p^*$ .
  - **Problem:** is  $p = \exp(h_{X,Y})$  always MTP<sub>2</sub> assuming that  $p^*$  is MTP<sub>2</sub>?
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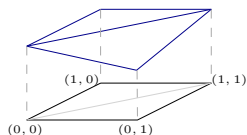
# When is the MLE the exponential of a tent function?

## Definition

Let  $X = \{x_1, \dots, x_n\} \subseteq \mathbb{R}^d$  be a *min-max closed* configuration. Then  $X$  is **tidy** if

The restriction of  $h_{X,y}$  to  $X$  is supermodular  $\iff$  The whole function  $h_{X,y}$  is supermodular.

## Example



If  $X = \{(0,0), (0,1), (1,0), (1,1)\}$ , then  $X$  is tidy because

$$y_{(0,0)} + y_{(1,1)} \geq y_{(0,1)} + y_{(1,0)} \implies h_{(X,y)} \text{ is supermodular.}$$

## Example

Consider again

$$X = \{(0,0,0), (6,0,0), (6,4,0), (8,4,2), (6,4,\frac{3}{2})\}.$$

- The restriction of *any*  $h_{X,y}$  to  $X$  is supermodular.
- But not all  $h_{X,y}$  are supermodular!  $\implies$  Not tidy.

# When is the MLE the exponential of a tent function?

## Theorem (R., Sturmfels, Tran, Uhler)

Let  $X \subseteq \mathbb{R}^d$  be min-max closed such that  $\text{conv}(X) = \text{MMconv}(X)$ .

Then,  $X$  is tidy if

- $X \subseteq \mathbb{R}^2$ , or
- $X \subseteq \{0, 1\}^d$ .

Therefore, the MLE for configurations in  $\mathbb{R}^2$  and in  $\{0, 1\}^d$  is always a tent function.

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## Conjecture

*These are the only tidy configurations.*

# Optimization Problem in the Tidy Case

## Theorem (R., Sturmfels, Tran, Uhler)

If  $X \subseteq \mathbb{R}^d$  is a tidy configuration, then,

- The MLE  $p^*$  is the exponential of a  $p^* = \exp(h_{X,y^*})$ , and
- The set of heights for which  $\exp(h_{X,y})$  is  $MTP_2$  is a convex polytope  $\mathcal{S}$ .

Therefore, we can use, e.g. projected gradient descent or the conditional gradient method, to find the best heights  $y^*$ .

$$\begin{array}{ll} \text{maximize}_y & \sum_{i=1}^n w_i y_i - \int \exp(h_{X,y}) \\ \text{s.t.} & y \in \mathcal{S}. \end{array}$$

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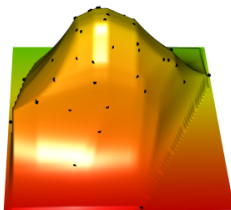
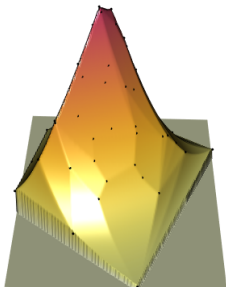
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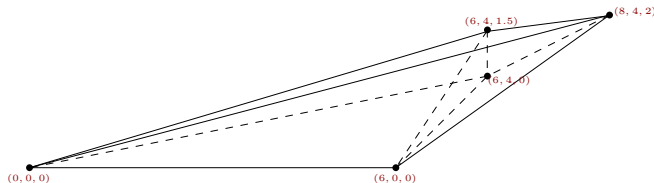
$$\begin{aligned} & \underset{y \in \mathcal{S}}{\text{maximize}_y} && \sum_{i=1}^n w_i y_i - \int \exp(h_{X,y}) \\ & \text{s.t.} && y \in \mathcal{S}. \end{aligned}$$





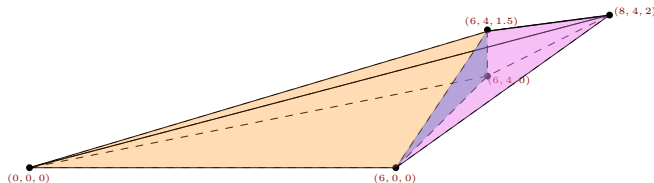
# What is the shape of the MLE in the general case?

- In  $\mathbb{R}^2$  and  $\{0, 1\}^d$  the MLE is the exponential of a tent function.
- If the log-concave MLE  $\phi$  is a supermodular tent function, then  $\phi$  is also the  $\text{MTP}_2$  log-concave MLE.
- Let  $X = \{(0, 0, 0), (6, 0, 0), (6, 4, 0), (8, 4, 2), (6, 4, \frac{3}{2})\}$ ,  $w = \frac{1}{28}(15, 1, 1, 1, 10)$ . The log-concave MLE  $\phi$  is not supermodular.



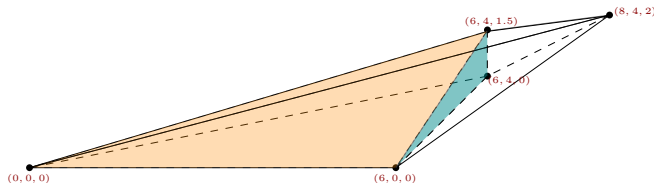
# What is the shape of the MLE in the general case?

- In  $\mathbb{R}^2$  and  $\{0, 1\}^d$  the MLE is the exponential of a tent function.
- If the log-concave MLE  $\phi$  is a bimonotone tent function, then  $\phi$  is also the  $\text{MTP}_2$  log-concave MLE.
- Let  $X = \{(0, 0, 0), (6, 0, 0), (6, 4, 0), (8, 4, 2), (6, 4, \frac{3}{2})\}$ ,  $w = \frac{1}{28}(15, 1, 1, 1, 10)$ . The log-concave MLE  $\phi$  is not bimonotone.



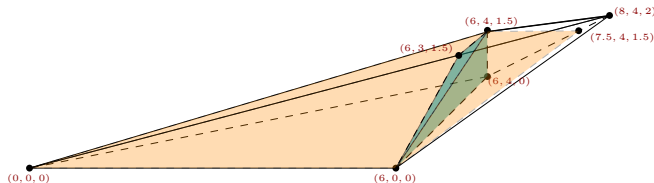
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the MLE is a tent function on  $X \cup \{(6,3,\frac{3}{2}), (7.5,4,\frac{3}{2})\}$  with subdivision as above.

## Conjecture

Let  $X = \{x_1, \dots, x_n\} \subset \mathbb{R}^d$  be a point configuration, and let  $w \in \mathbb{R}^n$  be the corresponding set of weights. Let  $\phi : \mathbb{R}^d \rightarrow \mathbb{R}$  be the log-concave maximum likelihood estimator (which is a tent function above  $X$ ), and let  $\Delta$  be the subdivision it induces.

1. If  $\Delta$  is a bimonotone subdivision, then  $\phi$  is also the  $MTP_2$  log-concave MLE.
2. If  $\Delta$  is not bimonotone, consider the hyperplanes spanned by each of the bimonotone codimension 1 cells of  $\Delta$ , and intersect  $\text{conv}(X)$  with them. Call this new subdivision  $\Delta'$ . The  $MTP_2$  log-concave maximum likelihood estimator is a piecewise linear function whose underlying subdivision is  $\Delta'$  or any subdivision refined by  $\Delta'$ .

# Summary and Remaining Questions

## Summary:

- We showed that the MLE under log-concavity and  $\text{MTP}_2$  exists and is unique with probability one.
- We showed that in some cases it is the exponential of a tent function, and we can compute it using convex optimization over a finite-dimensional convex set.
- We saw which tent functions are supermodular, i.e. are candidates for the MLE.

## Remaining questions and future work

- Characterize the shape of the MLE in the general case.
- Study the sample complexity of solving the problem.
- Design and analyze algorithms for finding the MLE.

# Announcement

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**9:30AM - 5PM**  
**MIT, E17-304**

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Thank you!